

# Adaptive Boxcar Deconvolution on Full Lebesgue Measure Sets\*

G erard Kerkyacharian, Dominique Picard  
*Universit e de Paris X-Nanterre and Universit es de Paris VI-VII*  
and Marc Raimondo  
*University of Sydney*

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## Abstract

We consider the non-parametric estimation of a function that is observed in white noise after convolution with a boxcar, the indicator of an interval  $(-a, a)$ . In a recent paper Johnstone et al. (2004) have developed a wavelet deconvolution algorithm (called **WaveD**) that can be used for “certain” boxcar kernels. For example, **WaveD** can be tuned to achieve near optimal rates over Besov spaces when  $a$  is a Badly Approximable (BA) irrational number. While the set of all BA’s contains quadratic irrationals e.g.  $a = \sqrt{5}$  it has Lebesgue measure zero, however. In this paper we derive two tuning scenarios of **WaveD** that are valid for “almost all” boxcar convolution (i.e. when  $a \in A$  where  $A$  is a full Lebesgue measure set). We propose (i) a tuning inspired from Minimax theory over Besov spaces; (ii) a tuning inspired from Maxiset theory providing similar rates as for BA numbers. Asymptotic theory informs that (i) in the worst case scenario, departure from the BA assumption, affects **WaveD** convergence rates, at most, by log factors; (ii) the Maxiset tuning, which yields smaller thresholds, is superior to the Minimax (conservative) tuning over a whole range of Besov sup-scales. Our asymptotic results are illustrated in an extensive simulation of boxcar convolution observed in white noise.

## 1 Introduction

We observe the stochastic process

$$Y_n(dt) = f \star b(t)dt + \sigma n^{-1/2}W(dt), \quad t \in T = [0, 1], \quad (1)$$

where  $b(t) = \frac{1}{2a}\mathbb{I}\{|t| \leq a\}$ ,  $\sigma$  is a positive constant,  $W(\cdot)$  is a Gaussian white noise and

$$f \star b(t) = \frac{1}{2a} \int_{-a}^a f(t-u)du. \quad (2)$$

This is an important model for the problem of recovery of noisy signals (or images) in linear motion blur, see Bertero and Boccacci (1998). Over the last decade, many wavelet methods have been developed to recover  $f$  from indirect observations Donoho (1995); Abramovich

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and Silverman (1998); Pensky and Vidakovic (1999); Walter and Shen (1999); Johnstone (1999); Fan and Koo (2002); Kalifa and Mallat (2003). However, the boxcar assumption (2) escapes most of the previously cited works. More recent papers which deals specifically with the boxcar problem includes Hall, Ruymgaart, van Gaans and van Rooij (2001), Neelamani, Choi and Baraniuk (2004), Johnstone and Raimondo (2004), and Johnstone, Kerkyacharian, Picard and Raimondo (2004) ([JKPR] in the sequel). The boxcar convolution has the special feature that if the boxcar half-width is rational then, certain frequencies are lost in the periodic model. If one changes the observation model as in Hall, Ruymgaart, van Gaans and van Rooij (2001) the latter difficulty disappear for signals  $f$  with compact support. Another approach to deal with the boxcar scenario is to introduce a regularisation step in the wavelet reconstruction algorithm see Neelamani, Choi and Baraniuk (2004). In this paper we follow the approach of [JKPR] and consider boxcar convolutions (2) where the boxcar half-width  $a$  is an irrational number. If  $a$  is chosen among Badly Approximable (BA) irrational numbers (those contains quadratic irrational like  $\sqrt{5}$ ) then the WaveD method is near optimal for a wide range of target functions and error losses. In the finite sample implementation  $(t_1, \dots, t_n)$  of the model (1) WaveD can recover the unknown function  $f$  with an accuracy of order

$$\left(\frac{\log n}{n}\right)^\beta, \quad \text{where} \quad (3)$$

$$\beta = \frac{sp}{4 + 2s}, \quad \text{if } s \geq \frac{2p}{\pi} - 2 \quad (4)$$

and

$$\beta = \frac{(s - 1/\pi + 1/p)p}{4 + 2(s - 1/\pi)}, \quad \text{if } \frac{1}{\pi} \leq s < \frac{2p}{\pi} - 2, \quad (5)$$

performance being measured in an integrated  $L^p$ -metric, for any  $p > 1$ . Here  $n$  denotes the usual sample size and  $s$  plays the role of a smoothness index for our target function  $f$ .

In fact near optimal properties of WaveD holds for any kernel function  $b_a(t)$  whose Fourier transforms  $(b_l(a))$  satisfies a decay condition when averaged over dyadic blocks, let

$$\tau_j^2(a) = |C_j|^{-1} \sum_{l=2^j}^{2^j+r} |b_l(a)|^{-2} \quad (6)$$

then for any kernel function  $b_a(t)$  such that  $\tau_j^2(a) \asymp 2^{3j}$ , the rate result (3) holds.

For statistical applications, an important issue is whether the WaveD estimator is robust against departures from the BA assumption. In this paper we combine the Maxiset theorem of Kerkyacharian and Picard (2000) with the equidistribution lemma of Johnstone and Raimondo (2004) to extend the results of [JKPR] outside the BA assumption. We propose two tuning scenarios of WaveD that can be applied to “almost all” boxcar convolutions (i.e.  $a \in A$  where  $A$  is a full Lebesgue measure set). In the first scenario we show that departure from the BA assumption affects rate (3) by, at most, log factors when tuning the WaveD algorithm over standard Besov spaces. In our second scenario we tune the WaveD algorithm to achieve rate (3) on certain Besov sub-scales. A theoretical comparison of the two scenarios suggests that smaller thresholds (such as arise in the second scenario) will always give better results than larger (conservative) thresholds (such as arise in the first scenario). This is confirmed by an extensive simulation study of boxcar convolution observed in white noise. All figures and tables presented in this paper can be reproduced using the WaveD1.3 software package available at <http://www.usyd.edu.au:8000/u/marcr/>.

We begin in section 2 by preliminaries on WaveD estimation, Besov spaces and Maxiset theorem. Asymptotic results are summarised in section 3 and numerical performances are studied in Section 4. Proofs are given in Section 5.

## 2 Preliminaries

### 2.1 Wavelet Deconvolution in a periodic setting

The wavelet deconvolution (WaveD) method proposed by [JKPR] combines both Fourier and Wavelet analysis. Let  $\Phi, \Psi$  denote the (periodised) Meyer scaling and wavelet function, see e.g. Meyer (1990), Mallat (1998). Let  $e_l(t) = e^{2\pi i l t}$ ,  $l \in \mathbb{Z}$  and write  $f_l = \langle f, e_l \rangle$ ,  $b_l = \langle b, e_l \rangle$  for the Fourier coefficients of  $f, b$  respectively where  $\langle f, g \rangle = \int_T f \bar{g}$ . The WaveD estimator is based on hard thresholding of a wavelet expansion as follows : (notice that here and in the sequel  $\kappa$  will denote the multiple index  $(j, k)$ ).

$$\hat{f}_n = \sum_{\kappa \in \Lambda_n} \hat{\beta}_\kappa \Psi_\kappa \mathbb{I}\{|\hat{\beta}_\kappa| \geq \hat{\sigma} \eta \sigma_j c_n\} \quad (7)$$

here and in the sequel  $\Psi_{-1} = \Phi$ . The wavelet coefficients are computed in the Fourier domain:

$$\hat{\beta}_\kappa = \sum_{l \in C_j} \left( \frac{y_l}{b_l} \right) \bar{\Psi}_l^\kappa \quad (8)$$

using eigen values of the boxcar function:

$$b_l = \frac{\sin \pi a l}{\pi a l}, l \in \mathbf{Z}. \quad (9)$$

Noting that for irrational number  $a$ , there are no zeros in (9), we use the fast algorithm of Donoho and Raimondo (2004) to compute the wavelet transform (8) and its inverse (7). This algorithm takes full advantage of the compact support of the Meyer wavelet in the Fourier domain:  $C_j = \{l : \Psi_l^\kappa \neq 0\} \subset (2\pi/3) \cdot [-2^{j+2}, -2^j] \cup [2^j, 2^{j+2}]$ .

The tuning parameters of WaveD are:

- The range of resolution levels (frequencies) where the approximation (7) is used:

$$\Lambda_n = \{(j, k), -1 \leq j \leq j_1, 0 \leq k \leq 2^j\},$$

here  $j_1$  determines the highest resolution level of WaveD. Theoretical properties of  $j_1$  are given in Section 3. In software, the default value of the finest scale  $j_1$  is determined from the data:  $j_1$  is set to be the level preceding  $j(100\%)$  where  $j(100\%)$  is the smallest level where 100% of thresholding occurs.

- The threshold value has four input parameters  $\hat{\sigma} \eta \sigma_j c_n$ 
  - $\hat{\sigma}$ : an estimate of the noise standard deviation,  $\sigma$ . If  $y_{J,k} = \langle Y_n, \Psi_{J,k} \rangle$ , denote the finest scale wavelet coefficients of the observed data, then  $\hat{\sigma} = m.a.d.\{y_{J,k}\}/.6745$ , where *m.a.d.* is median absolute deviation.
  - $\eta$ : a constant which depends on the tail of the Noise distribution. For Gaussian noise, the range  $\sqrt{2} \leq \eta \leq \sqrt{6}$  gives good result in practice. In software, the default value is  $\sqrt{6}$ . Theoretical properties of  $\eta$  are given in Section 3.

- $\sigma_j$ : is a level-dependent scaling factor based on the convolution kernel. Theoretical properties of  $\sigma_j$  are given in Sections 2 and 3. In software the (standard) default value is

$$\sigma_j^2 := \tau_j^2(a) = |C_j|^{-1} \sum_{l \in C_j} |b_l(a)|^{-2}$$

- $c_n$ : is a sample size-dependent scaling factor. Theoretical properties of  $c_n$  are given in Section 3. In software the default value is

$$c_n = \left( \frac{\log n}{n} \right)^{1/2}$$

## 2.2 Besov spaces of periodic functions

Let us first introduce the standard Besov spaces of periodic functions  $B_{\pi,r}^s(T)$ ,  $s > 0, \pi \geq 1$  and  $r \geq 1$ . For this purpose, define for every measurable function  $f$

$$\Delta_\varepsilon f(x) = f(x + \varepsilon) - f(x),$$

then, recursively,  $\Delta_\varepsilon 2f(x) = \Delta_\varepsilon(\Delta_\varepsilon f)(x)$  and similarly  $\Delta_\varepsilon^N f(x)$  for positive integer  $N$ . Let

$$\rho^N(t, f, \pi) = \sup_{|\varepsilon| \leq t} \left( \int_0^1 |\Delta_\varepsilon^N f(u)|^\pi du \right)^{1/\pi}.$$

Then for  $N > s$ , we define :

$$B_{\pi,r}^s(T) = \{f \text{ periodic} : \left( \int_0^1 \left( \frac{\rho^N(t, f, \pi)}{t^s} \right)^r \frac{dt}{t} \right)^{1/r} < \infty\}.$$

(with the usual modifications for  $r$  or  $\pi = \infty$ .)

In this setting, recall that the Besov spaces are characterised by the behaviour of the wavelet coefficients (as soon as the wavelet is periodic and has enough smoothness and vanishing moments).

**Definition 1.** For  $f \in L^\pi(T)$ ,

$$f = \sum_{j,k} \beta_{j,k} \Psi_{j,k} \in B_{\pi,r}^s(T) \iff \sum_{j \geq 0} 2^{j(s+1/2-1/\pi)r} \left[ \sum_{0 \leq k \leq 2^j} |\beta_{j,k}|^\pi \right]^{r/\pi} < \infty. \quad (10)$$

The Besov spaces have proved to be an interesting scale for studying the properties of statistical procedures. The index  $s$  indicates the degree of smoothness of the function. Due to the differential averaging effects of the integration parameters  $\pi$  and  $r$ , the Besov spaces capture a variety of smoothness features in a function including spatially inhomogeneous behaviour, see Donoho et al. (1995).

In order to fully describe asymptotic properties of WaveD for boxcar deconvolution it is useful to introduce the following Besov sub-scales:

**Definition 2.** For  $s > 0$ ,  $1 \leq \pi \leq \infty$ ,  $\tau \in \mathbf{R}$ , we define,

$$\tilde{B}_{\pi,\infty}^{s,\tau} = \left\{ f = \sum_{j \geq -1, k} \beta_{j,k} \psi_{j,k}, \sup |j|^\tau 2^{j(s+\frac{1}{2}-\frac{1}{\pi})} \|(\beta_{j,k})_k\|_{l_\pi} < \infty \right\} \quad (11)$$

The latter range of Besov scales are embedded into standard Besov scales.

$$\tilde{B}_{\pi,\infty}^{s,\tau} \subset B_{\pi,\infty}^s, \forall \tau \geq 0$$

### 2.3 The Maxiset-approach

The following theorem is borrowed from Kerkyacharian and Picard (2000). We refer to the appendix for condition (52) (known as the Temlyakov property). First, we introduce some notation:  $\mu$  will denote the measure such that for  $j \in \mathbf{N}$ ,  $k \in \mathbf{N}$ ,

$$\mu\{(j, k)\} = \|\sigma_j \Psi_{j,k}\|_p^p = \sigma_j^p 2^{j(\frac{p}{2}-1)} \|\Psi\|_p^p \quad (12)$$

$$l_{q,\infty}(\mu) = \left\{ f, \sup_{\lambda>0} \lambda^q \mu\{(j, k) / |\beta_{j,k}| > \sigma_j \lambda\} < \infty \right\} \quad (13)$$

**Theorem 1.** *Let  $p > 1$ ,  $0 < q < p$ ,  $\{\psi_{j,k}, j \geq -1, k = 0, 1, \dots, 2^j\}$  be a periodised wavelet basis of  $L_2(T)$  and  $\sigma_j$  be a positive sequence such that the heteroscedastic basis  $\sigma_j \psi_{j,k}$  satisfies property (52). Suppose that  $\Lambda_n$  is a set of pairs  $(j, k)$  and  $c_n$  is a deterministic sequence tending to zero with*

$$\sup_n \mu\{\Lambda_n\} c_n^p < \infty. \quad (14)$$

If for any  $n$  and any pair  $\kappa = (j, k) \in \Lambda_n$ , we have

$$\mathbf{E}|\hat{\beta}_\kappa - \beta_\kappa|^{2p} \leq C (\sigma_j c_n)^{2p} \quad (15)$$

$$P\left(|\hat{\beta}_\kappa - \beta_\kappa| \geq \eta \sigma_j c_n / 2\right) \leq C (c_n^{2p} \wedge c_n^4) \quad (16)$$

for some positive constants  $\eta$  and  $C$  then, the wavelet based estimator

$$\hat{f}_n = \sum_{\kappa \in \Lambda_n} \hat{\beta}_\kappa \psi_\kappa \mathbf{I}\{|\hat{\beta}_\kappa| \geq \eta \sigma_j c_n\} \quad (17)$$

is such that, for all positive integers  $n$ ,

$$\mathbf{E}\|\hat{f}_n - f\|_p^p \leq C c_n^{p-q},$$

if and only if :

$$f \in l_{q,\infty}(\mu), \quad \text{and}, \quad (18)$$

$$\sup_n c_n^{q-p} \left\| f - \sum_{\kappa \in \Lambda_n} \beta_\kappa \psi_\kappa \right\|_p^p < \infty. \quad (19)$$

This theorem identifies the 'Maxiset' of a general wavelet estimator of the form (17), by conditions (18) and (19). In [JKPR] p.565, we see that estimated wavelet coefficients ( $\hat{\beta}_\kappa$ ) defined at (8) are unbiased, normally distributed with variance bounded by the quantity  $\tau_j^2(a)$  defined at (6). Clearly conditions (15) and (16) heavily rely on the precise evaluation of this quantity. In the next section we show that for *almost all*  $a$ ,  $\tau_j^2(a) = O(2^{3j} j^{11(1+\delta)})$ . Hence, for such  $\tau_j$ , the proof arguments of [JKPR] which hold for  $\tau_j \asymp 2^{3j}$  no longer apply. In fact an improvement of Theorem 1 is required to derive the asymptotic theory for *almost all*  $a$ . The following corollary is an adaptation of the proof of theorem 1, Kerkyacharian and Picard (2000)

**Corollary 1.** *Let  $0 < q < \infty$ ,  $-\infty < \alpha < \infty$ . Let  $\xi(t) = \xi_{(q,\alpha)}(t)$  a continuous non decreasing function, such that  $\xi(0) = 0$  :*

$$\xi(t) = \begin{cases} t^q (\log(\frac{1}{t}))^\alpha, & t \in [0, \kappa] \\ (\log(\frac{1}{\kappa}))^\alpha t^q & t > \kappa \end{cases} \quad (20)$$

where  $0 < \kappa \leq \exp -\alpha/q$  if  $\alpha \geq 0$  ; and  $0 < \kappa < 1$  if  $\alpha < 0$ . Under the same hypothesis as in theorem 1 , the estimator  $\hat{f}_n$  is such that, for all positive integers  $n$ ,

$$\mathbf{E} \|\hat{f}_n - f\|_p^p \leq C \frac{c_n^p}{\xi(c_n)}$$

if and only if :

$$f \in l_{\xi, \infty}(\mu), \quad \text{and}, \quad (21)$$

$$\sup_n \frac{c_n^p}{\xi(c_n)} \left\| f - \sum_{\kappa \in \Lambda_n} \beta_\kappa \psi_\kappa \right\|_p^p < \infty. \quad (22)$$

where

$$l_{\xi, \infty}(\mu) = \left\{ f, \sup_{\lambda > 0} \xi(\lambda) \mu\{(j, k) / |\beta_{j,k}| > \sigma_j \lambda\} < \infty \right\} \quad (23)$$

*Remark 0.* Corollary 1 offers more flexibility than Theorem 1, in particular, it allows us to deal with scaling factor  $\sigma_j$  of the form  $\sigma_j \asymp 2^{\nu_j} j^z$ . While this has direct applications in the problem at hand, we note that there are interesting applications of Corollary 1 to multichannel deconvolution Pensky and Zayed (2002), see also the discussion by De Canditiis and Pensky in [JKPR].

## 2.4 Diophantine approximations

To every real number correspond a unique sequence  $(a_k)$ :

$$a = [a_0; a_1, a_2, \dots] = a_0 + \frac{1}{a_1 + \frac{1}{a_2 + \frac{1}{a_3 + \dots}}}. \quad (24)$$

For rational numbers the above expansion stops  $a = [a_0; a_1, a_2, \dots, a_k]$  and  $a_k = 0$  for later  $k$ , whereas for irrational number  $a$  the sequence of  $(a_k)$ ,  $a_k > 0$ , is infinite. The rational numbers defined by terminating the expansion (24) at stage  $k: p_k(a)/q_k(a) = [a_0, a_1, \dots, a_k]$  are called the *convergents* of  $a$ . For any irrational  $a$ , the convergents have the property of *best approximation*: for  $n \geq 1$ ,

$$\inf_{1 \leq k \leq q_n} \|ka\| = |q_n a - p_n| = \|q_n a\|, \quad (25)$$

where  $\|x\|$  denotes the distance from  $x \in \mathbb{R}$  to the nearest integer. The study of such Diophantine approximations plays a central role in our analysis of the boxcar blur, since from (9):  $b_l = \sin(\pi l a)/(\pi l a)$  from which it follows that

$$\frac{2}{\pi} \frac{\|la\|}{la} \leq b_l \leq \frac{\|la\|}{la}. \quad (26)$$

Hence, the properties of WaveD (7) depends on the nature of the irrational number  $a$ .

**Definition 3.** An irrational number  $a$  is called *Badly Approximable (BA)* if

$$\sup_n \frac{q_n}{q_{n-1}} < \infty.$$

For boxcar with BA widths, [JKPR] have shown that the WaveD estimator achieves near optimal rates of convergence over Besov spaces. Another interesting class of irrational numbers can be derived from the “measure theory” of continued fractions Khintchine (1963)

**Definition 4.** For all  $a \in A_\delta$ , where  $A_\delta$  is a full Lebesgue measure set, we have that for all  $n \geq n_0$ ,

$$\frac{q_n}{q_{n-1}} < (\log q_{n-1})^{(1+\delta)}. \quad (27)$$

In the sequel “almost all  $a$ ” means “for all  $a$  in  $A_\delta$ ”.

One of the main difficulties of boxcar deconvolution is that Fourier coefficients  $b_l(a), l = 0, 1, 2, \dots$  vary erratically according to the approximations  $\|la\|/l, l = 1, 2, \dots$ , see Figure 1. This difficulty disappears when averaging over dyadic blocks, let

$$\tau_j^2(a) = |C_j|^{-1} \sum_{l \in C_j} |b_l(a)|^{-2} \quad (28)$$

where  $C_j = \{l : \Psi_l^\kappa \neq 0\} \subset (2\pi/3) \cdot [-2^{j+2}, -2^j] \cup [2^j, 2^{j+2}]$ , then:

**Proposition 1.** Let  $(b_l)$  be the Fourier coefficients of the boxcar kernel  $b(t) = \frac{1}{2a} \mathbf{I}\{|t| \leq a\}$ .

1. For BA boxcar scale  $a$

$$\tau_j^2(a) \leq c_1 2^{3j}. \quad (29)$$

2. For each  $\delta > 0$ , there is a constant  $c_1 > 0$  such that for almost all  $a$ :

$$\tau_j^2(a) \leq c_1 2^{3j} j^z, \quad z = 11(1 + \delta). \quad (30)$$

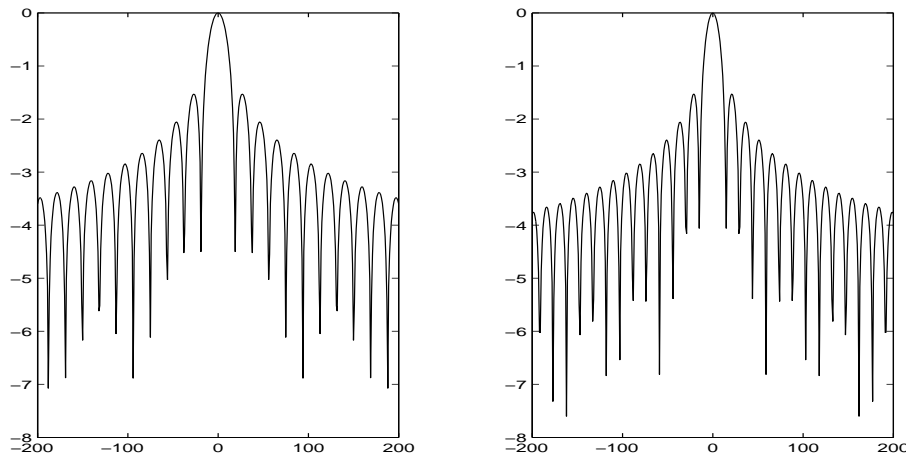


Figure 1: Eigen values of the boxcar function with: (left) BA scale  $a = 1/\sqrt{353}$ ; (right)  $a$  randomly chosen from a  $\mathcal{U}(0.025, 0.075)$  distribution.

*Remark 1.* The bound (29) was derived in [JKPR] and is presented here for comparison purposes. The novelty appears in bound (30) where we see that for almost all scale  $a$  the scaling factor  $\tau_j(a)$  is the same as in the BA case up-to log terms. In the next section we study the effect of these extra log terms on the asymptotic properties of WaveD.

### 3 Asymptotic results

For boxcar convolutions observed in white noise we derive tuning parameters  $\Lambda_n, \eta, \sigma_j$  and  $c_n$  which yield near optimal properties of WaveD for *almost all*  $a$ . Using the Maxiset approach we present two possible tuning scenarios of the waved estimator. Our first scenario (section 3.1) is inspired from the Minimax theory: we fix a smoothness class (here  $B_{\pi,r}^s$ ) and find a set of tuning parameters which ensure near-optimal asymptotic properties of WaveD over the entire smoothness class. In our second scenario (section 3.2) the primary focus is the convergence rate rather than the smoothness class: we propose a set of tuning parameters with which WaveD achieves the rate (3) known to be achievable for boxcar convolution with Badly Approximable scale. Finally, we give a theoretical comparison of the two scenarios in section 3.3.

#### 3.1 Tuning inspired from Minimax theory

To study the properties of WaveD over standard Besov spaces one must find conditions under which a particular Besov space  $B_{\pi,r}^s$  may be embedded in the spaces  $l_{q,\infty}(\mu)$  as well as imply the condition (19). We follow the steps of [JKPR] (Appendix B.1) taking into account the Degree of Ill Posedness (DIP)  $\nu = 3/2$  of the Boxcar convolution Johnstone and Raimondo (2004). Recalling that the Maxisets  $l_{q,\infty}(\mu)$  are defined in terms of the measure  $\mu\{(j, k)\} = c\sigma_j^p 2^{j(\frac{p}{2}-1)}$ ; a key condition in recovering the standard Besov spaces is to let

$$\sigma_j^2 =: \sigma_j^2(1) = c_1 2^{3j} \quad (31)$$

Letting as usual  $\Lambda_n = \{(j, k), -1 \leq j \leq j_1, 0 \leq k \leq 2^j\}$  and  $z = 11(1 + \delta)$  we compensate for the extra log-term appearing in (30) by taking

$$c_n =: c_n(1) = \left(\frac{\log n}{n}\right)^{1/2} (\log n)^{\frac{z}{2}} \text{ and } 2^{j_1} = n^{\frac{1}{4}} (\log n)^{-\frac{1+z}{4}}. \quad (32)$$

These choices ensure that conditions (15) and (16) are satisfied.

**Proposition 2.** *Suppose that we observe the random process (1) with  $\sigma = 1$ . Let  $p > 1$  be an arbitrary number. If  $f$  belongs to  $B_{\pi,r}^s(T)$  with  $\pi \geq 1$ ,  $s \geq 1/\pi$  and  $0 < r \leq \frac{2p-1}{2+s-1/\pi}$  if  $s + 2 = \frac{2p}{\pi}$ ,  $0 < r \leq \infty$  otherwise, then, for  $\eta \geq 2\sqrt{8\pi(p \vee 2)}$ ,  $z = 11(1 + \delta)$  the WaveD estimator (7) with tuning parameters (31), (32) and  $\hat{\sigma} = 1$ , is such that, for almost all  $a$ :*

$$\mathbf{E}\|\hat{f}_1 - f\|_p^p \leq C[n^{-1}(\log(n))^{(1+z)\beta}]^\beta, \text{ for all positive integers } n, \quad (33)$$

where  $\beta$  depends on  $s$  as shown at (4) and (5).

*Remarks 2.* A comparison of the rates (3) and (33) shows that, for almost all boxcar convolutions (in the sense of definition 2), the rate properties of WaveD are, in the worst-case-scenario, only affected up to log factors. These results are consistent with those derived in Johnstone and Raimondo (2004) regarding the Degree of Ill Posedness of the boxcar deconvolution.

#### 3.2 Tuning inspired from Maxiset theory

Corollary 1 allows a certain degree of freedom in the tuning of WaveD. More specifically, moment condition (15) and tail behavior (16) may be satisfied for a different choice of  $\sigma_j, c_n$



than that proposed in section 3.1. For example, it is possible to fit WaveD with a slightly smaller threshold than in proposition 2 and yet derive a consistent estimator. While this second scenario may not ensure optimal properties over the scale standard Besov spaces, it is interesting to note that it yields faster rate of convergence over Besov sub-scales  $\tilde{B}_{\pi,r}^{s,\tau}$ . In fact, it is possible to tune WaveD in such a way that it achieves the rate (3) as available for boxcar convolution with BA scales. Following [JKPR] (Appendix A.1), we start by letting:

$$c_n =: c_n(2) = \left(\frac{\log n}{n}\right)^{1/2} \quad \text{and} \quad 2^{j_1} = n^{\frac{1}{4}}(\log n)^{-\frac{1+z}{4}}. \quad (34)$$

Now, let  $z = 11(1 + \delta)$  we compensate for the extra log terms in (30) by taking

$$\sigma_j^2 =: \sigma^2(2) = c_1 2^{3j} j^z. \quad (35)$$

**Proposition 3.** *Suppose that we observe the random process (1) with  $\sigma = 1$ . Let  $p > 1$  be an arbitrary number. If  $f$  belongs to  $\tilde{B}_{\pi,\infty}^{s,\tau}(T)$  with  $\pi \geq 1$ ,  $s \geq 1/\pi$  and with  $\tau \geq \frac{zs}{4}$  if  $s + 2 > \frac{2p}{\pi}$ ,  $\tau > 1 + \frac{zs}{4}$  if  $s + 2 = \frac{2p}{\pi}$ ,  $\tau \in \mathbf{R}$  otherwise, then, for  $\eta \geq 2\sqrt{8\pi(p \vee 2)}$ ,  $z = 11(1 + \delta)$  the WaveD estimator (7) with tuning parameters (35), (34) and  $\hat{\sigma} = 1$ , is such that, for almost all  $a$ :*

$$\mathbf{E}\|\hat{f}_2 - f\|_p^p \leq C[n^{-1} \log(n)]^\beta, \quad \text{for all positive integers } n, \quad (36)$$

where  $\beta$  is given at (4),(5).

*Remarks 3.* A quick look at Propositions 2 and 3 shows that they do not give a fair comparison between the two scenarios. The reason is that both the spaces and the convergence rates with which each tuning method is prescribed are different. Since Besov sub-scales are embedded in standard Besov spaces ( $\tilde{B}_{\pi,\infty}^{s,\tau} \subset B_{\pi,\infty}^s$ ) it is not surprising to observe better rate performances in the second scenario. A further application of corollary 1 allows a fair comparison of the two scenarios (next section).

### 3.3 Minimax versus Maxisets

The following proposition describes into details the differences between the two tuning scenarios over Besov sub-scales in terms of their respective Maxisets.

**Proposition 4.** *Suppose that we observe the random process (1) with  $\sigma = 1$ . Let  $p > 1$  be an arbitrary number, and  $q$  such that  $0 < q < p$ . Let  $v_n = (n^{-1} \log n)^{\frac{p-q}{2}}$ ,  $j_1 \geq 0$  chosen as in (32), then for*

$$\mathcal{F} = \left\{ f : \sup_n v_n \|f - \sum_{\kappa,j \leq j_1} \beta_\kappa \psi_\kappa\|_p^p < \infty \right\} \quad (37)$$

Denote  $\hat{f}_1$  be the WaveD estimator (7) with tuning parameters (31), (32) and  $\hat{f}_2$  the WaveD estimator (7) with tuning parameters (35), (34). Let  $\hat{\sigma} = 1$ , then, for almost all  $a$ :

(a)

$$\mathbf{E}\|\hat{f}_1 - f\|_p^p \leq C v_n, \quad \text{for all positive integers } n,$$

if and only if :

$$f \in MAX(1) = \{f \in l_{\xi_{(q, \frac{z(p-q)}{2})}, \infty}(\mu) \cap \mathcal{F}\}$$

for

$$l_{\xi_{(q, \frac{z(p-q)}{2})}, \infty}(\mu) = \left\{ f, \sup_{\lambda > 0} \xi_{(q, \frac{z(p-q)}{2})}(\lambda) \mu\{(j, k) : |\beta_{j,k}| > \sigma_j(1)\lambda\} < \infty \right\} \quad (38)$$

$$\mu\{(j, k)\} = 2^{j(2p-1)}.$$

(b)

$$\mathbf{E}\|\hat{f}_2 - f\|_p^p \leq Cv_n, \text{ for all positive integers } n, \quad (39)$$

if and only if :

$$f \in MAX(2) = \{f \in l_{q, \infty}(\tilde{\mu}) \cap \mathcal{F}\}.$$

for

$$f \in l_{q, \infty}(\tilde{\mu}) = \left\{ f, \sup_{\lambda > 0} \lambda^q \tilde{\mu}\{(j, k) : |\beta_{j,k}| > \sigma_j(2)\lambda\} < \infty \right\}, \quad (40)$$

$$\tilde{\mu}\{(j, k)\} = 2^{j(2p-1)} j^{\frac{zp}{2}}.$$

*Remarks 4.* Proposition 4 shows that over Besov sub-scales both tuning scenarios yield similar rate of convergence (3). The difference between the two methods appears in their respective Maxisets described by conditions (38) and (40). Hence the two maxisets to be compared are:  $MAX(1)$  for minimax-WaveD and  $MAX(2)$  for maxiset-WaveD.

**Proposition 5.** *Under the same assumptions as in proposition 4, we have*

$$MAX(1) \subset MAX(2)$$

*Remark 5.* This result shows that if we compare the tuning scenarios using the maxiset point of view, the second tuning (Maxiset) is always better than the first tuning (Minimax) since it achieves a near optimal rate of convergence over a larger class of functions. This suggests that the smaller threshold setting (34), (35) may give better result in practice than the larger threshold setting (31),(32). This is confirmed by our simulation study (section 5).

## 4 Numerical performances

We study the finite sample properties of the WaveD algorithm (7) when applied to noisy boxcar convolution (1). Figure 1 depicts four inhomogeneous signals borrowed from the statistical literature Johnstone et al. (2004), Donoho and Raimondo (2004). Figure 2 depicts the signals of Figure 1 after blurring with a boxcar kernel. In Figure 3, we added Gaussian white noise to each signals of Figure 2 (with medium noise level). In our simulation study we used three noise levels: low, medium and high as seen table 1. Our main results (summarised in section 3) state that the WaveD estimator can be applied to noisy boxcar convolution where the boxcar scale  $a$  is chosen randomly with respect to a continuous distribution. For illustration purposes, we used the Uniform (0.025, 0.075) distribution to set the boxcar parameter in each simulation of the model (1), as illustrated is in Figures 2 and 3. In the table below we give Monte Carlo approximation to the Root Mean Integrated Square Error (RMISE) of the WaveD estimator when fitted with the Minimax tuning (32), (31) and when fitted with the Maxiset tuning (34), (35). For comparison purposes we included the results

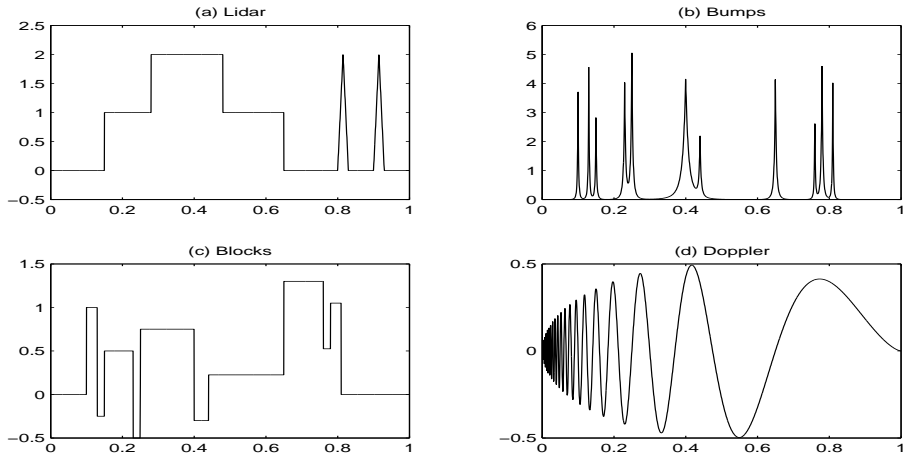


Figure 2: Four (inhomogeneous) signals,  $t_i = i/n$ ,  $n = 2048$ .

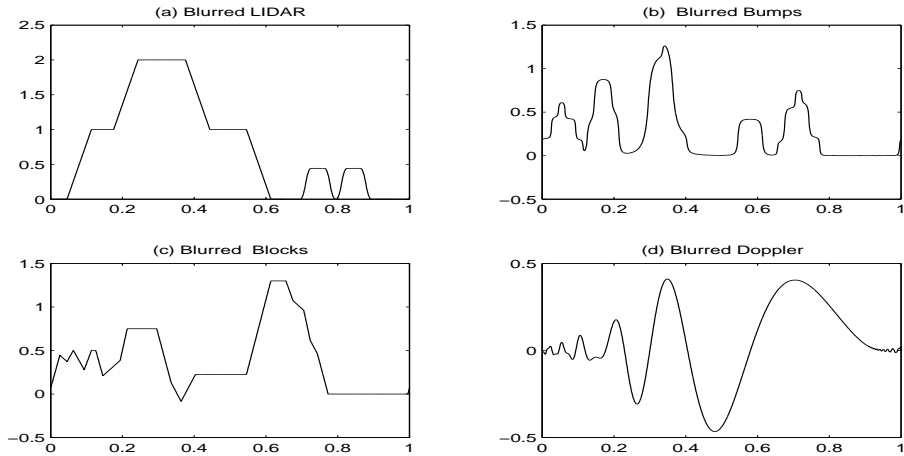


Figure 3: Signals of figure 1 after boxcar blurring with  $a$  randomly chosen from a  $\mathcal{U}(0.025, 0.75)$  distribution.

of standard WaveD [JKRP] when applied to a noisy boxcar convolution when the boxcar scale is  $a$  is a BA number ( $a = 1/\sqrt{353}$ ) which in the table below is indicated by boxcar-scale=BA. For the Minimax and Maxiset tunings we simulated noisy boxcar convolution with  $a$  randomly chosen from a  $\mathcal{U}(0.025, 0.75)$  distribution which in the table below is indicated by boxcar-scale=AA ('Almost All').

*Analysis of the results.* Our numerical study confirms the theoretical analysis of section 3. First, we see that the WaveD estimator can be applied successfully to Almost All boxcar convolutions. The results obtained with smaller thresholds (as with Maxiset) are better than those obtained with larger thresholds (as with Minimax). Finally, comparing the results obtained in the BA case with those of the AA setting we note slightly poorer performances in the AA case, as to be expected Johnstone and Raimondo (2004). The WaveD estimator is based on Hard Thresholding and enjoys fast computation Donoho and Raimondo (2004). Alternative thresholding strategies such as block-wise thresholding Pensky and Vidakovic (1999); Cavalier and Tsybakov (2002), multichannel approach Pensky and Zayed (2002)

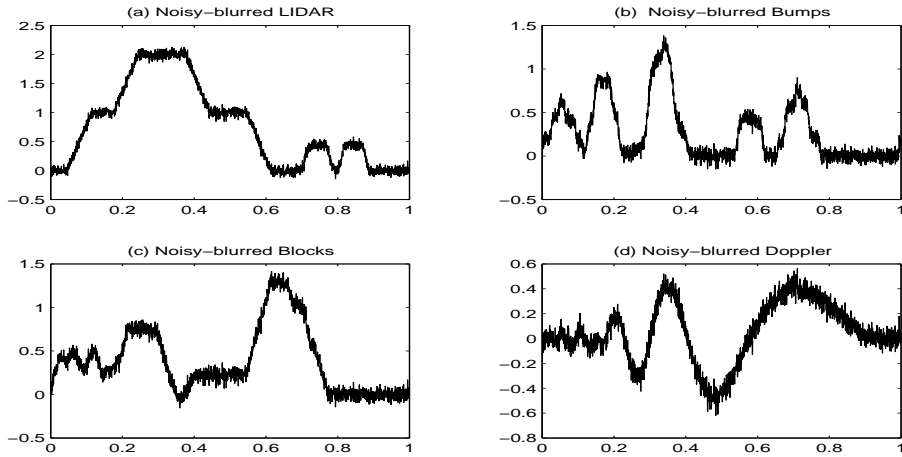


Figure 4: Illustration of the model (1) with signals of Figure 2, Gaussian noise with  $\sigma = 0.05$ .

Table 1: Monte-Carlo approximations to  $RMISE = \sqrt{\mathbf{E}\|\hat{f} - f\|_2^2}$ . The results are means of 1000 independent simulations of the model (1) with  $n = 2048$  illustrated as Figure 3.

Tuning	Boxcar-scale	Signal	$\sigma_{low} = 0.005$	$\sigma_{med} = 0.05$	$\sigma_{high} = 0.1$
Standard	BA	Lidar	0.0990	0.2084	0.2744
Maxiset	AA	Lidar	0.1220	0.2671	0.3400
Minimax	AA	Lidar	0.1379	0.3274	0.3863
Standard	BA	Bumps	0.2042	0.4563	0.5233
Maxiset	AA	Bumps	0.2933	0.5103	0.5466
Minimax	AA	Bumps	0.3231	0.5332	0.5749
Standard	BA	Blocks	0.1207	0.2287	0.2676
Maxiset	AA	Blocks	0.1469	0.2643	0.3097
Minimax	AA	Blocks	0.1626	0.3044	0.3324
Standard	BA	Doppler	0.0601	0.1063	0.1372
Maxiset	AA	Doppler	0.0681	0.1346	0.1605
Minimax	AA	Doppler	0.0754	0.1572	0.1872

may also give good results.

## 5 Proofs

*Proof of Proposition 1.* We start by a lemma deducible from Johnstone and Raimondo (2004).

**Lemma 1.** *Let  $p/q$  and  $p'/q'$  be successive principal convergents in the continued fraction expansion of a real number  $a$  and let  $A_\delta$  be a full Lebesgue measure set. Let  $q \geq 4$  and  $N$  be a non-negative integer with  $N + q < q'$ . Then, for all numbers  $a \in A_\delta$ ,*

$$\sum_{l=N+1}^{N+q} \|la\|^{-2} \leq c_1(a)q^2 (\log q)^{2+2\delta} \quad (41)$$

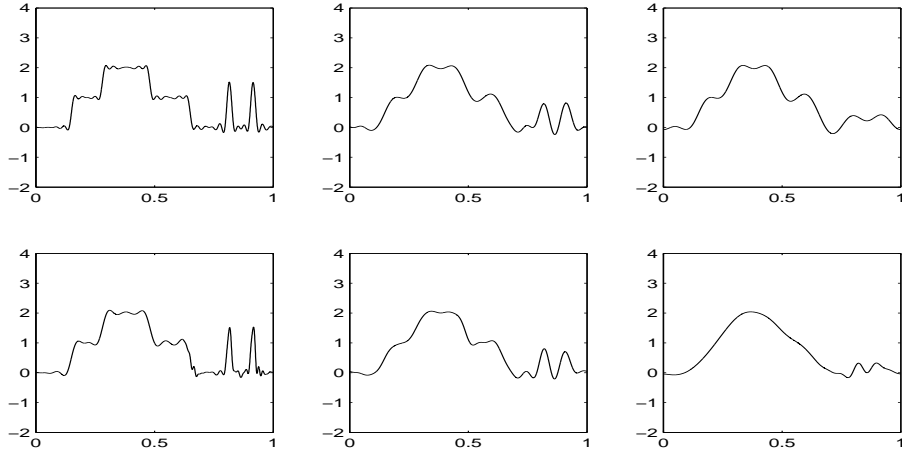


Figure 5: WaveD Lidar estimates in low-med-high noise level (left to right). Top plots: data drawn from a noisy boxcar convolution with a BA scale. Bottom plots: data drawn from a noisy boxcar convolution with  $a$  randomly chosen from a  $\mathcal{U}(0.025, 0.075)$  distribution.

By definition of the Meyer wavelet  $C_j \subset \{l : 2^j \leq |l| \leq 2^{j+2}\}$ . To simplify the exposition we will use the symmetry of  $x \rightarrow x2$  about 0, and consider that from now on that  $C_j \subset \{l : 2^j \leq l \leq 2^{j+2}\}$ . Let  $m$  be the smallest integer such that  $q_m \geq 2^j$ . From the geometric growth of the convergents  $q_{m+2r} \geq 2^r q_m$ , hence:

$$C_j \subset \mathbf{N} \cap [q_{m-1}, q_{m+4}).$$

Introducing sets  $D_\tau = \mathbf{N} \cap [q_{m+\tau-1}, q_{m+\tau})$ ,  $\tau = 0, 1, \dots, 4$  whose union covers  $C_j$ . For all  $a \in A_\delta$  we have that  $q_{m+\tau} \leq q_{m+\tau-1} (\log q_{m+\tau-1})^{1+\delta}$ . Hence, there are at most  $K_\tau = (\log q_{m+\tau-1})^{1+\delta}$  blocks of length  $q_{m+\tau-1}$  to cover  $D_\tau$ ; we apply lemma 1 within each block:

$$\sum_{D_\tau} \|la\|^{-2} = \sum_{\text{blocks}} \sum_{l \in \text{block}} \|la\|^{-2} \leq K_\tau c_1 (q_{m+\tau-1})^2 K_\tau 2 = c_1 (q_{m+\tau-1})^2 K_\tau 3.$$

Taking log in (27) we see that there exists a constant  $c_2 = C(\delta, \tau)$  such that:

$$K_\tau \leq c_2 K_0, \quad q_{m+\tau-1} \leq c_2 q_{m-1} K_0^\tau, \quad \tau = 1, \dots, 4 \quad (42)$$

It follows that

$$\begin{aligned} \sum_{C_j} \|la\|^{-2} &\leq \sum_{\tau=0}^4 \sum_{D_\tau} \|la\|^{-2} \leq c_1 (q_{m-1}^2 K_0 3 + q_m^2 K_1 3 + q_{m+1}^2 K_1 3 + q_{m+2}^2 K_3 3 + q_{m+3}^2 K_\tau 4) \\ &\leq c_2 K_0 3 q_{m-1}^2 (1 + K_0 2 + K_0 4 + K_0 6 + K_0 8) \leq c_3 K_0^{11} q_{m-1}^2. \end{aligned}$$

By construction  $q_{m-1} \leq 2^j$  and so  $K_0 = (\log q_{m-1})^{1+\delta} \leq c_4 j^{1+\delta}$  which combined with (26) shows that

$$\tau_j^2 = |C_j|^{-1} \sum_{l \in C_j} |b_l|^{-2} \leq C 2^j \sum_{C_j} \|la\|^{-2} \leq C 2^{3j} j^{11(1+\delta)}$$

□

*Proof of Proposition 2.* We consider the WaveD estimator (7) with the Minimax tuning (31), (32) and we apply Theorem 1. In this setting, arguments similar to those in [JKPR] (Appendix A) are used to prove that the following claims hold (C1): Inequalities (15) and (16) hold with  $\eta \geq 2\sqrt{8\pi(p \vee 2)}$ . (C2): the basis  $(\sigma_j \Psi_{jk})$  satisfies condition (52). (C3): Conditions (14) is satisfied with our choice of parameters. Hence, Theorem 1 applies to (7) which gives the following maxisets/rate result:

$$\mathbf{E} \|\hat{f}_1 - f\|_p^p \leq C c_n^{p-q},$$

if and only if :

$$f \in l_{q,\infty}(\mu) = \left\{ f, \sup_{\lambda>0} \lambda^q \mu\{(j, k) : |\beta_{jk}| > \sigma_j \lambda\} < \infty \right\}, \quad \text{and}, \quad (43)$$

$$\sup_n c_n^{q-p} \left\| f - \sum_{\kappa \in \Lambda_n} \beta_\kappa \psi_\kappa \right\|_p^p < \infty. \quad (44)$$

for

$$\mu\{(j, k)\} = 2^{j(2p-1)}. \quad (45)$$

We take  $p - q = 2\beta$  and derive the rate results (33) for standard Besov spaces. Condition (44) essentially uses the same arguments as in [JKPR] (Appendix B.1).

For condition (47), as in [JKPR], we suppose that  $f \in B_{\pi,r}^s(T)$  and distinguish between the cases  $s + 2 \geq \frac{2p}{\pi}$  and  $s + 2 < \frac{2p}{\pi}$ . In the first case, we take  $q = \frac{2p}{s+2}$  and have (all the inequalities below are true up to obvious absolute constants):

$$\begin{aligned} \mu\{(j, k) : |\beta_{jk}| > 2^{3j/2} \lambda\} &= \sum_{jk} 2^{j(2p-1)} I\{|\beta_{jk}| > 2^{3j/2} \lambda\} \\ &\leq \sum_j 2^{2pj} \wedge 2^{j(2p-1)} \sum_k \left[ \frac{|\beta_{jk}|}{2^{3j/2} \lambda} \right]^\pi \\ &\leq \sum_j 2^{2pj} \wedge \left[ \frac{1}{\lambda} \right]^\pi 2^{-j(s\pi+2(\pi-p))} \varepsilon_j^\pi \end{aligned} \quad (46)$$

where  $(\varepsilon_j)$  in a sequence belonging to  $l_r$ . Let  $J$  such that  $2^{-J(s+2)} \sim \lambda$ , we get:

$$\begin{aligned} \mu\{(j, k) : |\beta_{jk}| > 2^{3j/2} \lambda\} &\leq \sum_{j \leq J} 2^{2pj} + \sum_{j > J} \left[ \frac{1}{\lambda} \right]^\pi 2^{-j(s\pi+2(\pi-p))} \varepsilon_j^\pi \\ &\leq \lambda^{\frac{-2p}{s+2}} + \left[ \frac{1}{\lambda} \right]^\pi \left[ \lambda^{\frac{\pi(s+2)-2p}{s+2}} I\left\{s + 2 > \frac{2p}{\pi}\right\} + \sum_j \varepsilon_j^\pi I\left\{s + 2 = \frac{2p}{\pi}\right\} \right] \end{aligned}$$

Which ends the proof of this case. Now, if  $s + 2 < \frac{2p}{\pi}$ , we take  $q = \frac{2p-1}{2+s-1/\pi}$ , and use Sobolev embeddings:  $B_{\pi,r}^s(T) \subset B_{q,r}^{s'}(T)$  for  $s' - 1/q = s - 1/\pi$ . Using (46) with  $q$  instead of  $\pi$ , we get:

$$\begin{aligned} \mu\{(j, k) : |\beta_{jk}| > 2^{3j/2} \lambda\} &\leq \lambda^{-q} \sum_j 2^{-j(s'q+2(q-p))} \varepsilon_j^q \\ &\leq \lambda^{-q} \sum_j 2^{-j \frac{2p-\pi(s+2)}{2+s-1/\pi}} \end{aligned}$$

□

*Proof of Proposition 3.* We consider the WaveD estimator (7) with the Maxiset tuning (34), (35) and apply we apply Theorem 1. In this setting, arguments similar to those in [JKPR] (Appendix A) are used to prove that the following claims hold (C1): Inequalities (15) and (16) hold with  $\eta \geq 2\sqrt{8\pi(p \vee 2)}$ . (C2): the basis  $(\sigma_j \Psi_{jk})$  satisfies condition (52). (C3): Conditions (14) is satisfied with our choice of parameters. Hence, Theorem 1 applies to (7) which gives the following maxisets/rate result:

$$\mathbf{E} \|\hat{f}_2 - f\|_p^p \leq C c_n^{p-q},$$

if and only if :

$$f \in l_{q,\infty}(\tilde{\mu}) = \left\{ f, \sup_{\lambda>0} \lambda^q \tilde{\mu}\{(j,k) : |\beta_{j,k}| > \sigma_j(2)\lambda\} < \infty \right\}, \quad \text{and,} \quad (47)$$

$$\sup_n c_n^{q-p} \left\| f - \sum_{\kappa \in \Lambda_n} \beta_\kappa \psi_\kappa \right\|_p^p < \infty.$$

for

$$\tilde{\mu}\{(j,k)\} = 2^{j(2p-1)} j^{\frac{zp}{2}}. \quad (48)$$

Using standard wavelet arguments, it is elementary using the definitions of  $j_1$  and  $c_n$  to see that

$$\sup_n c_n^{q-p} \left\| f - \sum_{\kappa \in \Lambda_n} \beta_\kappa \psi_\kappa \right\|_p^p < \infty \iff \sup_n c_n^{q-p} \sum_{j \geq j_1} 2^{j(\frac{p}{2}-1)} \sum_k |\beta_{jk}|^p < \infty$$

And this is obviously true if  $f \in \tilde{B}_{p\infty}^{\frac{2(p-q)}{p}, \frac{z(p-q)}{2p}}$ . Now, again we take  $\beta = p - q$  and prove rate result (36) over Besov sub-scales (2), using arguments quite analogous to those of Proposition 2. First we observe that the inclusion into the space  $B_{p\infty}^{\frac{2(p-q)}{p}, \frac{z(p-q)}{2p}}$  is similar to the proof given in [JKPR].

For the inclusion into  $l_{q,\infty}(\tilde{\mu})$ , we have as in (46), in the case,  $s + 2 \geq \frac{2p}{\pi}$ ,

$$\begin{aligned} \tilde{\mu}\{(j,k) : |\beta_{jk}| > 2^{3j/2} j^{z/2} \lambda\} &= \sum_{jk} 2^{j(2p-1)} j^{zp/2} I\{|\beta_{jk}| > 2^{3j/2} j^{z/2} \lambda\} \\ &\leq \sum_j 2^{2pj} j^{zp/2} \wedge 2^{j(2p-1)} j^{zp/2} \sum_k \left[ \frac{|\beta_{jk}|}{2^{3j/2} j^{z/2} \lambda} \right]^\pi \\ &\leq \sum_j 2^{2pj} j^{zp/2} \wedge \left[ \frac{1}{\lambda} \right]^\pi 2^{-j(s\pi+2(\pi-p))} j^{\frac{z(p-\pi)}{2} - \tau\pi} \end{aligned}$$

Let  $J$  such that  $2^J \sim \lambda^{\frac{1}{s+2}} \log \frac{1}{\lambda}^{-z/4}$ . We get :

$$\begin{aligned} \tilde{\mu}\{(j,k) : |\beta_{jk}| > 2^{3j/2} j^{z/2} \lambda\} &\leq \sum_{j \leq J} 2^{2pj} j^{zp/2} + \sum_{j > J} \left[ \frac{1}{\lambda} \right]^\pi 2^{-j(s\pi+2(\pi-p))} j^{\frac{z(p-\pi)}{2} - \tau\pi} \\ &\leq \lambda^{\frac{-2p}{s+2}} + \left[ \frac{1}{\lambda} \right]^\pi \left[ \lambda^{\frac{\pi(s+2)-2p}{s+2}} \log \frac{1}{\lambda}^{\frac{zs\pi}{4} - \tau\pi} I\{s+2 > \frac{2p}{\pi}\} \right] \\ &+ \sum_j j^{\frac{zs\pi}{4} - \tau\pi} I\{s+2 = \frac{2p}{\pi}\} \end{aligned}$$

Which ends the proof of this case. Now, again if  $s + 2 < \frac{2p}{\pi}$ , we take  $q = \frac{2p-1}{2+s-1/\pi}$  and use Sobolev embeddings and (46) with  $q$  instead of  $\pi$ , we get:

$$\begin{aligned} \tilde{\mu}\{(j, k) : |\beta_{jk}| > 2^{3j/2} j^{z/2} \lambda\} &\leq \lambda^{-q} \sum_j 2^{-j(s'q+2(q-p))} j^{\frac{z(p-q)}{2}-\tau q} \\ &\leq \lambda^{-q} \sum_j 2^{-j\frac{2p-\pi(s+2)}{2+s-1/\pi}} j^{\frac{z(p-q)}{2}-\tau q} \end{aligned}$$

□

*Proof of Proposition 4.* It is a consequence of Theorem 1 and corollary 1.

*Proof of proposition 5.*

**Lemma 2.** Let  $0 < q < \infty$ ,  $-\infty < \alpha < \infty$ . Let us define  $\xi(t) = \xi_{q,\alpha}(t)$  a continuous non decreasing function, such that  $\xi(0) = 0$  :

$$\xi(t) = \begin{cases} t^q (\log(\frac{1}{t}))^\alpha, & t \in [0, \kappa] \\ (\log(\frac{1}{\kappa}))^\alpha t^q & t > \kappa \end{cases}$$

where  $0 < \kappa \leq \exp -\alpha/q$  if  $\alpha \geq 0$  ; and  $0 < \kappa < 1$  if  $\alpha < 0$ . Then

$$\exists C_0, \quad \forall \lambda > 0, \quad \sum_{j \geq 0} \frac{1}{\xi(2^j \lambda)} \leq \frac{C_0}{\xi(\lambda)} \quad (49)$$

$$\forall p > q, \quad \exists C_p, \quad \forall \lambda > 0, \quad \sum_{j \geq 0} \frac{1}{2^{jp} \xi(2^{-j} \lambda)} \leq \frac{C_p}{\xi(\lambda)} \quad (50)$$

### Proof

1. Proof of (49) :

Let  $j_0 = \inf\{j \in \mathbf{N}, 2^j \lambda > \kappa\}$

$$\sum_{j \geq 0} \frac{1}{\xi(2^j \lambda)} = \sum_{j < j_0} \frac{1}{(2^j \lambda)^q (\log 1/2^j \lambda)^\alpha} + \sum_{j \geq j_0} \frac{1}{(\log(\frac{1}{\kappa}))^\alpha (2^j \lambda)^q}$$

The result is clearly obvious if  $j_0 = 0$ . Now if  $j_0 > 0$ , and  $0 \leq j < j_0$ , then  $\lambda \leq 2^j \lambda \leq \kappa < 2^{j_0} \lambda < 2\kappa$ ,

$$\sum_{j \geq 0} \frac{1}{\xi(2^j \lambda)} \leq \frac{1}{(\log 1/\lambda)^\alpha \lambda^q} \sum_{j < j_0} 2^{-jq} \left(\frac{\log 1/\lambda}{\log 1/2^j \lambda}\right)^\alpha + c_q \frac{1}{(\log(\frac{1}{\kappa}))^\alpha (2^{j_0} \lambda)^q}$$

But as  $(\log 1/\lambda)^\alpha \lambda^q$  is a non decreasing function on  $[0, \kappa]$ ,

$$\frac{1}{(\log(\frac{1}{\kappa}))^\alpha (2^{j_0} \lambda)^q} \leq \frac{1}{(\log(\frac{1}{\kappa}))^\alpha \kappa^q} \leq \frac{1}{(\log 1/\lambda)^\alpha \lambda^q}$$

and , for  $0 \leq j < j_0$

$$1 \leq \frac{\log 1/\lambda}{\log 1/2^j \lambda} = \frac{\log 1/2^j \lambda + j \log 2}{\log 1/2^j \lambda} = 1 + j \frac{\log 2}{\log 1/2^j \lambda} \leq 1 + j \frac{\log 2}{\log 1/\kappa}$$



and if  $\alpha \geq 0$ ,

$$\sum_{j < j_0} 2^{-jq} \left( \frac{\log 1/\lambda}{\log 1/2^j \lambda} \right)^\alpha \leq \sum_{j < j_0} 2^{-jq} \left( 1 + j \frac{\log 2}{\log 1/\kappa} \right)^\alpha \leq C$$

If  $\alpha < 0$

$$\sum_{j < j_0} 2^{-jq} \left( \frac{\log 1/\lambda}{\log 1/2^j \lambda} \right)^\alpha \leq \sum_{j < j_0} 2^{-jq} \leq C$$

2. Proof of (50) : Let  $j_0 = \inf\{j \in \mathbf{N}, 2^{-j}\lambda \leq \kappa\}$ ,  $j_0 = 0$  if  $\lambda \leq \kappa$ .

$$\begin{aligned} \sum_{j \geq 0} \frac{1}{2^{jp} \xi(2^{-j}\lambda)} &= \sum_{j \geq j_0} \frac{1}{2^{jp} (2^{-j}\lambda)^q (\log(\frac{1}{2^{-j}\lambda}))^\alpha} + \sum_{0 \leq j < j_0} \frac{1}{(\log(\frac{1}{\kappa}))^\alpha 2^{jp} (2^{-j}\lambda)^q} \\ &= \frac{1}{(\log 1/\lambda)^\alpha \lambda^q} \sum_{j \geq j_0} 2^{-j(p-q)} \left( \frac{\log(\frac{1}{\lambda})}{\log(\frac{1}{2^{-j}\lambda})} \right)^\alpha + \frac{1}{(\log(\frac{1}{\kappa}))^\alpha \lambda^q} \sum_{0 \leq j < j_0} 2^{-j(p-q)} \end{aligned}$$

If  $\lambda \leq \kappa$  we have only the first term and

$$\text{if } \alpha \geq 0, \quad \left( \frac{\log(\frac{1}{\lambda})}{\log(\frac{1}{2^{-j}\lambda})} \right)^\alpha \leq 1$$

$$\text{if } \alpha < 0, \quad \left( \frac{\log(\frac{1}{\lambda})}{\log(\frac{1}{2^{-j}\lambda})} \right)^\alpha \leq \left( 1 + \frac{j \log 2}{\log(\frac{1}{\kappa})} \right)^{|\alpha|}$$

and we get the result.

If now  $j_0 < \infty$ , for  $2^{-j_0}\lambda \leq \kappa < 2^{-(j_0-1)}\lambda$

$$\text{if } \alpha \geq 0, \quad \frac{1}{\lambda^q} \sum_{j \geq j_0} \frac{2^{-j(p-q)}}{(\log(\frac{1}{2^{-j}\lambda}))^\alpha} \leq \left( \frac{1}{\log 1/\kappa} \right)^\alpha \frac{2^{-j_0(p-q)}}{\lambda^q} \sum_{j \geq j_0} 2^{-(j-j_0)(p-q)} \leq \frac{C}{\lambda^p}$$

$$\text{if } \alpha < 0, \quad \frac{1}{\lambda^q} \sum_{j \geq j_0} \frac{2^{-j(p-q)}}{(\log(\frac{1}{2^{-j}\lambda}))^\alpha} \leq \frac{2^{-j_0(p-q)}}{\lambda^q} \sum_{j \geq j_0} 2^{-(j-j_0)(p-q)} \left( \log(\frac{1}{2^{-j_0}\lambda} + (j-j_0)) \right)^{|\alpha|} \leq \frac{C}{\lambda^p}$$

But as  $\lambda \geq \kappa$ ,  $\frac{1}{\lambda^p} = O(\frac{1}{\lambda^q})$

**Corollary 2.** Let  $I$  be a set,  $q < p$ , and  $\forall i, \mu(i) \geq 0$ . For all  $\gamma : I \mapsto \mathbf{R}$  we have the following equivalence :

1.

$$\exists C > 0, \quad \forall \lambda > 0, \quad \sum_{\{i, |\gamma(i)| > \lambda\}} \mu(i) \leq \frac{C}{\xi(\lambda)}$$

2.

$$\exists C', \quad \forall \lambda > 0, \quad \sum_{\{i, |\gamma(i)| \leq \lambda\}} |\gamma(i)|^p \mu(i) \leq C' \frac{\lambda^p}{\xi(\lambda)}$$

**Proof**

1  $\implies$  2 : We use (50)

$$\begin{aligned} \sum_{\{i, |\gamma(i)| \leq \lambda\}} |\gamma(i)|^p \mu(i) &= \sum_{j \in \mathbf{N}} \sum_{\{i, 2^{-j-1} < |\gamma(i)| \leq 2^{-j}\lambda\}} |\gamma(i)|^p \mu(i) \leq \\ \lambda^p \sum_{j \in \mathbf{N}} 2^{-jp} \sum_{\{i, 2^{-j-1} < |\gamma(i)|\}} \mu(i) &\leq \lambda^p \sum_{j \in \mathbf{N}} 2^{-jp} \frac{C}{\xi(2^{-j-1}\lambda)} \leq 2CC_p \frac{\lambda^p}{\xi(\lambda)} \end{aligned}$$

2  $\implies$  1 : We use (49)

$$\begin{aligned} \sum_{\{i, |\gamma(i)| > \lambda\}} \mu(i) &= \sum_{j \in \mathbf{N}} \sum_{\{i, 2^{j+1}\lambda \geq |\gamma(i)| > 2^j\lambda\}} \mu(i) \leq \sum_{j \in \mathbf{N}} \frac{1}{2^{jp}\lambda^p} \sum_{\{i, 2^{j+1}\lambda \geq |\gamma(i)|\}} \mu(i) |\gamma(i)|^p \\ &\leq \sum_{j \in \mathbf{N}} \frac{1}{2^{jp}\lambda^p} C' \frac{2^{(j+1)p}\lambda^p}{\xi(2^{j+1}\lambda)} \leq 2^p C' C_0 \frac{1}{\xi(\lambda)} \end{aligned}$$

□

Applying the previous lemma to  $\mu(j, k) = 2^{j(p/2-1)}\tau_j^p$  and  $\gamma(jk) = \frac{\beta_{jk}}{\tau_j}$  we derive the following lemma.

**Lemma 3.** *If  $\tau_j$  an arbitrary positive sequence,  $p > q$ ,  $\beta$  are positive real numbers, the following assertions are equivalent:*

1. *There exists  $C$ , such that:*

$$\forall \lambda > 0, \sum_{j \geq -1, k} 2^{j(\frac{p}{2}-1)} \tau_j^p I\{|\frac{\beta_{jk}}{\tau_j}| \geq \lambda\} \leq C \frac{1}{\xi_{(q, \beta)}(\lambda)}$$

2. *There exists  $C_1$ , such that:*

$$\forall \lambda > 0, \sum_{j, k} |\beta_{jk}|^p I\{|\frac{\beta_{jk}}{\tau_j}| \leq \lambda\} 2^{j(\frac{p}{2}-1)} \leq C_p \frac{\lambda^p}{\xi_{(q, \beta)}(\lambda)}$$

To prove proposition 5, let us suppose that  $f \in MAX(1)$ , and prove that  $f \in l_{q, \infty}(\tilde{\mu})$ . Let  $C$  be a generic constant which may change from line to line. Let us observe first that

$$\forall n \in \mathbf{N}, [n^{-1} \log(n)]^{\frac{q-p}{2}} \|f - \sum_{\kappa, j \leq j_1} \beta_{\kappa} \psi_{\kappa}\|_p^p \leq C < \infty, \text{ for } 2^{j_1} = n^{\frac{1}{4}} (\log n)^{-\frac{1+z}{4}},$$

which implies that

$$\forall j_1 \in \mathbf{N}, \|f - \sum_{\kappa, j \leq j_1} \beta_{\kappa} \psi_{\kappa}\|_p \leq C 2^{\frac{-2(p-q)j_1}{p}} j_1^{\frac{-z(p-q)}{p}}.$$

$$\forall j \in \mathbf{N}, (\sum_k |\beta_{jk}|^p 2^{j(\frac{1}{2}-\frac{1}{p})})^{1/p} \leq C 2^{\frac{-2(p-q)j}{p}} j^{\frac{-z(p-q)}{p}},$$

and more generally

$$\forall J > j \in \mathbf{N}, \sum_{j < j' < J, k} |\beta_{j'k}|^p 2^{j'(\frac{p}{2}-1)} \leq C 2^{-2(p-q)j} j^{-z(p-q)}. \quad (51)$$

To apply lemma 3 we have to check that

$$\sum_{jk} |\beta_{jk}|^p 2^{j(\frac{p}{2}-1)} 1_{|\frac{\beta_{jk}}{\sigma_j(2)}| \leq \lambda} \leq \lambda^{p-q}.$$

Obviously, from (51) we have only to check this for  $\lambda$  in a neighbourhood of 0,  $[0, \kappa]$  say. Now, let us introduce  $j_0$  such that  $j_0 \sim \log \frac{1}{\lambda}$ . Let us consider:

$$\begin{aligned} \sum_{jk} |\beta_{jk}|^p 2^{j(\frac{p}{2}-1)} 1_{|\frac{\beta_{jk}}{\sigma_j(2)}| \leq \lambda} &= \left[ \sum_{j \leq j_0} + \sum_{j > j_0} \right] \sum_k |\beta_{jk}|^p 2^{j(\frac{p}{2}-1)} 1_{|\frac{\beta_{jk}}{\sigma_j(2)}| \leq \lambda} \\ &\leq \sum_{j \leq j_0} \sum_k |\beta_{jk}|^p 2^{j(\frac{p}{2}-1)} 1_{|\frac{\beta_{jk}}{\sigma_j(1)}| \leq \lambda (\log \frac{1}{\lambda})^{\frac{z}{2}}} + \sum_{j > j_0} \sum_k |\beta_{jk}|^p 2^{j(\frac{p}{2}-1)} \\ &\leq C \left[ \lambda (\log \frac{1}{\lambda})^{\frac{z}{2}} \right]^{p-q} (\log \frac{1}{\lambda})^{-\frac{z(p-q)}{2}} + C' 2^{-j_0 2(p-q)} \leq C'' \lambda^{p-q} \end{aligned}$$

The first term of the last inequality is obtained by another application of lemma 3 using the assumption  $f \in l_{\xi_{q, \frac{z(p-q)}{2}}, \infty}(\mu)$ . The second term is obtained using (51). □

*Appendix: Temlyakov inequalities.* Let us recall the Temlyakov property for a basis  $e_n(x)$  in  $L^p$ : there exists absolute constants  $c, C$  such that for all  $\Lambda \subset \mathbf{N}$ ,

$$c \sum_{n \in \Lambda} \int |e_n(x)|^p dx \leq \int \left( \sum_{n \in \Lambda} |e_n(x)|^2 \right)^{p/2} dx \leq C \sum_{n \in \Lambda} \int |e_n(x)|^p dx$$

or, equivalently :

$$c' \left\| \left( \sum_{n \in \Lambda} |e_n(x)|^p \right)^{1/p} \right\|_p \leq \left\| \left( \sum_{n \in \Lambda} |e_n(x)|^2 \right)^{1/2} \right\|_p \leq C' \left\| \left( \sum_{n \in \Lambda} |e_n(x)|^p \right)^{1/p} \right\|_p \quad (52)$$

## References

- Abramovich, F. and Silverman, B. (1998), ‘Wavelet decomposition approaches to statistical inverse problems’, *Biometrika* **85**(1), 115–129.
- Bertero, M. and Boccacci, P. (1998), *Introduction to Inverse Problems in Imaging*, Institute of Physics, Bristol and Philadelphia.
- Cavalier, L. and Tsybakov, A. (2002), ‘Sharp adaptation for inverse problems with random noise’, *Probab. Theory Related Fields* **123**(3), 323–354.
- Donoho, D. (1995), ‘Nonlinear solution of linear inverse problems by wavelet-vaguelette decomposition’, *Applied Computational and Harmonic Analysis* **2**, 101–126.
- Donoho, D. L., Johnstone, I. M., Kerkyacharian, G. and Picard, D. (1995), ‘Wavelet shrinkage: Asymptopia?’, *Journal of the Royal Statistical Society, Series B* **57**, 301–369. With Discussion.
- Donoho, D. and Raimondo, M. (2004), ‘Translation invariant deconvolution in a periodic setting’, *The International Journal of Wavelets, Multiresolution and Information Processing* **14**(1). In press.

- Fan, J. and Koo, J. (2002), ‘Wavelet deconvolution’, *IEEE Transactions on Information Theory* **48**(3), 734–747.
- Hall, P., Ruymgaart, F., van Gaans, O. and van Rooij, A. (2001), Inverting noisy integral equations using wavelet expansions: A class of irregular convolutions, in ‘State of the Art in Probability and Statistics: Festschrift for Willem R. van Zwet’, Vol. 36 of *Lecture Notes-Monograph Series*, Institute of Mathematical Statistics, pp. 533–546.
- Johnstone, I., Kerkycharian, G., Picard, D. and Raimondo, M. (2004), ‘Wavelet deconvolution in a periodic setting’, *Journal of the Royal Statistical Society, Series B* **66**(3), 547–573. with discussion pp.627-652.
- Johnstone, I. M. (1999), ‘Wavelet shrinkage for correlated data and inverse problems: adaptivity results’, *Statistica Sinica* **9**(1), 51–83.
- Johnstone, I. M. and Raimondo, M. (2004), ‘Periodic boxcar deconvolution and diophantine approximation’, *Annals of Statistics* **32**(5), 1781–1804.
- Kalifa, J. and Mallat, S. (2003), ‘Thresholding estimators for linear inverse problems and deconvolutions’, *Annals of Statistics* **31**, 58–109.
- Kerkycharian, G. and Picard, D. (2000), ‘Thresholding algorithms and well-concentrated bases’, *Test* **9**(2).
- Khintchine, A. (1963), *Continued Fractions*, Noordhoff, Groningen.
- Mallat, S. (1998), *A wavelet tour of signal processing (2nd Edition)*, Academic Press Inc., San Diego, CA.
- Meyer, Y. (1990), *Ondelettes et Opérateurs-I*, Hermann.
- Neelamani, R., Choi, H. and Baraniuk, R. (2004), ‘Forward: Fourier-wavelet regularized deconvolution for ill-conditioned systems’, *IEEE Transactions on signal processing* **52**, 418–433.
- Pensky, M. and Vidakovic, B. (1999), ‘Adaptive wavelet estimator for nonparametric density deconvolution’, *Annals of Statistics* **27**, 2033–2053.
- Pensky, M. and Zayed, A. (2002), ‘Density deconvolution of different conditional distributions’, *Ann. Inst. Statist. Math.* **54**(3), 701–712.
- Walter, G. and Shen, X. (1999), ‘Deconvolution using the meyer wavelet’, *Journal of Integral Equations and Applications* **11**, 515–534.